

# THE OPTIMAL PLANNED REPLACEMENT MODEL WITH MINIMAL AND PERFECT REPAIR

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## 1. INTRODUCTION

In this paper we investigate a replacement model with two types of repairs. Repairs are classified into minimal and perfect repair. An operating unit is completely replaced whenever it reaches age  $\tau$  ( $\tau > 0$ ) at a cost  $c_1$  (planned replacement). If it fails at age  $t < \tau$ , it is either restores by a entire unit with probability  $p(t)$  at a cost  $c_2$  (perfect repair), or it undergoes minimal repair with probability  $\bar{p}(t) = 1 - p(t)$  at a cost  $c_3$ . After a planned replacement, the procedure is repeated. The aim of this paper is to find a condition to exist the optimal planned replacement period  $\tau^*$  which minimizes the total long-run expected cost per unit time.

## 2. PROBLEM FORMULATION

Let  $X_k$  be the time interval failure time between  $(k - 1)$ st. If  $F$  is the failure time distribution of the unit, then the failure time distribution following minimal repair for a unit which fails at time  $t$  is given by

$$P\left[X_k \leq x \left| \sum_{i=1}^{k-1} X_i = t \right.\right] = \frac{F(t+x) - F(t)}{\bar{F}(t)}, \quad x > t.$$

Let  $t$  be the age of a unit and  $\lambda(t)$  be the failure rate (or the hazard rate) function belonging to  $F(t)$ . Let  $G(t)$  be the distribution function of random variable  $Y =$  total time among perfect repairs. Formally,  $Y_1, Y_2, \dots, Y_{n-1}$  be a sequence of non-negative independent random variables with a common distribution  $G(t)$ . Then

$$Y = \sum_{i=1}^{n-1} Y_i.$$

Also let  $G(t) = P(Y \leq t)$ .

The survival distribution of the time between successive perfect repairs is given by

$$(1) \quad \bar{G}(t) = \exp\left[-\int_0^t p(x)\lambda(x) dx\right],$$

where  $\bar{G}(t) = 1 - G(t)$  and  $G(0) = 0$ . Hence  $r(t) = p(t)\lambda(t)$  denotes the failure rate function belonging to  $G(t)$ .

Let  $\{\hat{N}(t), t \geq 0\}$  be the number of perfect repairs that occur during  $[0, t]$ .

Let  $\{N(t), t \geq 0\}$  be the number of minimal repairs that occur during  $[0, t]$ .

Let  $\{N^*(\tau), \tau \geq 0\}$  be the number of minimal repairs whenever it reaches age  $\tau$  from a final perfect repair.

Therefore, we constitute the total time  $W$  and the total cost  $C$  until a unit completely replaced whenever it reaches age  $\tau$  is given by

$$(2) \quad W = \sum_{i=1}^{n-1} Y_i + \tau,$$

if  $Y_1 < \tau, Y_2 < \tau, \dots, Y_{n-1} < \tau, Y_n \geq \tau$ .

And

$$(3) \quad C = c_1 + c_2 \hat{N}(W) + c_3 [N(W) + N^*(\tau)],$$

if  $Y_1 < \tau, Y_2 < \tau, \dots, Y_{n-1} < \tau, Y_n \geq \tau$ .

## 3. ANALYSIS OF OPTIMAL MODEL

In this section, we shall explain two part of expected total time  $E[W]$  and expected total cost  $E[C]$ .

**Lemma 3.1.**  $E[W]$  denotes the expected total time of the unit until planned replacement. Thus it follow that

$$E[W] = \frac{\int_0^\tau \bar{G}(t) dt}{\bar{G}(\tau)}.$$

**Lemma 3.2.** Let  $N^*(\tau)$  denotes the number of minimal repairs in  $[0, \min(Y_n, \tau)]$ , for all  $\tau > 0$ , where  $Y_n$  denotes the random length of the final perfect repair before planned replacement.

$$E[N^*(\tau)] = \Lambda(\tau) - R(\tau).$$

**Lemma 3.3.** Let  $N(W)$  denotes the random number of minimal repairs during  $[0, W]$ .

$$E[N(W)] = \frac{1}{\bar{G}(\tau)} \left[ \int_0^\tau \Lambda(t) dG(t) + [1 + R(\tau)]\bar{G}(\tau) - 1 \right].$$

**Lemma 3.4.**  $\hat{N}(W)$  be the random number of perfect repairs between  $[0, W]$ . If only perfect repair actions are taken, the renewal process is well known to be a process  $\{\hat{N}(W), W \geq 0\}$ .

$$E[\hat{N}(W)] = \frac{G(\tau)}{\bar{G}(\tau)}.$$

**Theorem 3.1.** The total long-run expected cost per unit time can be obtained by using the theory of renewal process and is equal to:

$$(4) \quad K(\tau) = \frac{c_1 + c_2 E[\hat{N}(W)] + c_3 E[N(W) + N^*(\tau)]}{E[W]}.$$

**Theorem 3.2.** Let  $F(t)$  have failure rate function  $\lambda(t)$  and suppose that the functions  $r(t)$  is continuous. Then if  $\lambda$  and  $r$  are monotonically increasing function and monotonically decreasing function, respectively. And that

$$\lim_{x \rightarrow +\infty} \lambda(x) \int_0^x \bar{G}(t) dt > \frac{c_1}{c_3},$$

there exists at least one finite positive period  $\tau^*$  which minimizes the total long-run expected cost per unit time  $K(\tau)$ .

*Proof.* We can differentiate  $K$  with respect to  $\tau$ . Noting that  $d\bar{G}(\tau)/d\tau = -p(\tau)\lambda(\tau)\bar{G}(\tau)$  and  $dR(\tau)/d\tau = r(\tau) = p(\tau)\lambda(\tau)$ . From (4), it follows that

$$(5) \quad \frac{dK(\tau)}{d\tau} = \frac{\bar{G}(\tau)}{[\int_0^\tau \bar{G}(t) dt]^2} \Theta(\tau) = 0,$$

where

$$\Theta(\tau) = c_3 \left[ \int_0^\tau [\lambda(\tau) - \Lambda(t)r(t)] \bar{G}(t) dt - \bar{G}(\tau)\Lambda(\tau) \right] - [c_1 + c_3 - c_2] \left[ \int_0^\tau [r(\tau) - r(t)] \bar{G}(t) dt \right] - c_1. \quad (6)$$

We assumed that the cost of planned replacement is higher than the cost of perfect repair and minimal repair. Differentiating the right-hand side in (6), we have

$$c_3 \left[ \lambda'(\tau) \int_0^\tau \bar{G}(t) dt \right] - [c_1 + c_3 - c_2] \left[ r'(\tau) \int_0^\tau \bar{G}(t) dt \right],$$

which is nonnegative, if we assume  $\lambda'(\tau) \geq 0$  and  $r'(\tau) \leq 0$ .

In (6) is a continuous increasing function of  $\tau$  which is negative ( $-c_1$ ) at  $\tau \rightarrow 0$  and  $c_1 > c_2 \geq c_3$ . And tends to  $+\infty$  if  $\lim_{\tau \rightarrow +\infty} \lambda(t) = \lambda(+\infty) = +\infty$  as  $\tau \rightarrow +\infty$ . Hence there always exists a unique solution  $\tau = \tau^*$ ,  $0 < \tau^* < +\infty$  of (5). Since  $dK(\tau)/d\tau$  has the same pattern  $(-\infty, 0, +\infty)$ , it follows that  $K(\tau)$  has a minimizes at  $\tau^*$ . Under the strictly increasing assumption,  $K(\tau)$  is strictly increasing, so  $\tau^*$  is unique. If  $\tau^*$  is the solution, then from (4) and (6) it is easy to get

$$(7) \quad K(\tau^*) = c_3 \lambda(\tau^*) - (c_1 + c_3 - c_2) r(\tau^*).$$

**Remark 3.1.** If  $c_1 + c_3 < c_2$ , the optimal planned replacement age is  $\tau^* = +\infty$ .

#### 4. EXAMPLE

In this section, we consider the particular case where  $p(t) \equiv p$ ,  $0 < p < 1$ . In this case, we assume that the probability of a perfect repair does

not depend on age. We have  $r(t) = p\lambda(t)$ ,  $G(t) = 1 - [F(t)]^p$ , and

$$(8) \quad K_p(\tau) = \frac{c_1 \bar{G}(\tau) + \left[ c_2 + \left( \frac{1-p}{p} \right) c_3 \right] G(\tau)}{\int_0^\tau \bar{G}(t) dt}.$$

The p.d.f. of the Weibull distribution with shape parameter  $\alpha$  and scale parameter  $\beta$ .

From (8), it follows that

$$(9) \quad \frac{dK_p(\tau)}{d\tau} = \frac{\bar{G}(\tau)}{[\int_0^\tau \bar{G}(t) dt]^2} \theta(\tau) = 0,$$

where

$$\theta(\tau) = \left[ c_2 + \left( \frac{1-p}{p} \right) c_3 - c_1 \right] \left[ \lambda(\tau) \int_0^\tau \bar{G}(t) dt + \frac{1}{p} \bar{G}(\tau) \right] - \frac{1}{p} \left[ c_2 + \left( \frac{1-p}{p} \right) c_3 \right].$$

Consider the case in which  $c_2 + [(1-p)/p]c_3 \leq c_1$ ,  $\theta(\tau) < 0$  for  $\tau > 0$ , then we have the optimal planned replacement period is  $\tau = +\infty$ . Supposing that the  $c_2 + [(1-p)/p]c_3 > c_1$ , and that

$$\lim_{x \rightarrow +\infty} \lambda(x) \int_0^x \bar{G}(t) dt > \frac{1}{p} \left[ \frac{c_2 + \left( \frac{1-p}{p} \right) c_3}{c_2 + \left( \frac{1-p}{p} \right) c_3 - c_1} \right].$$

By (9), If  $\alpha > 1$ , a finite optimal planned replacement period  $\tau^*$  exists, since  $\lambda(+\infty) = +\infty$ .  $\theta$  is increasing since  $\lambda(\tau)$  and easy computation shows that

$$\theta'(\tau) = \lambda'(\tau) \int_0^\tau \bar{G}(t) dt \geq 0.$$

Also note that  $\theta(0) = -c_1/p < 0$ . Thus if the  $\theta(\tau) = 0$  has a solution  $\tau^*$ , then  $K'_p(\tau)$  is negative on  $(0, \tau^*)$  and positive on  $(\tau^*, +\infty)$ .

$$\lambda(\tau^*) \int_0^{\tau^*} \bar{G}(t) dt - G(\tau^*) = \frac{pc_1}{p(c_2 - c_1) + (1-p)c_3}.$$

Equation (9) reduces to the equation

$$\left( \frac{\alpha}{\beta} \right) \left( \frac{t}{\beta} \right)^{\alpha-1} \int_0^\tau \exp \left[ -p \left( \frac{t}{\beta} \right)^\alpha \right] dt + \frac{1}{p} \exp \left[ -p \left( \frac{t}{\beta} \right)^\alpha \right] = \frac{1}{p} \left[ \frac{c_2 + \left( \frac{1-p}{p} \right) c_3}{c_2 + \left( \frac{1-p}{p} \right) c_3 - c_1} \right],$$

which is easily solved by numerical methods. If  $\alpha = 1$ , the equation above has no solution.

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